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Wolves make roadways safer, generating large economic returns to predator conservation

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Abstract

Recent studies uncover cascading ecological effects resulting from removing and reintroducing predators into a landscape, but little is known about effects on human lives and property. We quantify the effects of restoring wolf populations by evaluating their influence on deer-vehicle collisions (DVCs) in Wisconsin. We show that, for the average county, wolf entry reduced DVCs by 24%, yielding an economic benefit that is 63 times greater than the costs of verified wolf predation on livestock. Most of the reduction is due to a behavioral response of deer to wolves, rather than through a deer population decline from wolf predation. This finding supports ecological research emphasizing the role of predators in creating a “landscape of fear.” It suggests wolves control economic damages from overabundant deer in ways that human deer hunters cannot.

Significance Statement

Measuring the economic benefits conveyed by predators is difficult—often, effects are indirect and operate through complex ecological changes. As a result, debates about the expansion of predators has pit salient costs against more speculative estimates of benefits that might be dismissed as unreliable or ideologically motivated. We quantify the indirect benefits of wolves (*Canis lupus*) to human lives and property through reductions in deer-vehicle collisions. Moreover, we decompose the effect into two components: changes in prey behavior versus prey abundance. This decomposition is important when effective policy depends on whether hunters can replicate the effects of predators. In the case of wolves, we conclude that human deer hunters cannot.

Main Text

Introduction

Populations of apex predators have declined across the world’s landscapes over the past two hundred years due to government bounty programs, hunting pressure, habitat loss, and declines in prey populations (1, 2). Ecologists are beginning to unravel the far-reaching ecological effects of these changes (1–5), but little is known about the economic effects. While many of the costs attributable to predators are salient and quantifiable, such as predation on livestock and pets, estimating and valuing the often subtle and indirect beneficial effects of predators is more difficult (6, 7).¹

The recent expansion of the gray wolf (*Canis lupus*) offers a unique opportunity to concretely measure the cascading benefits of a predator whose reintroduction is controversial. Wolves once ranged over most of the northern hemisphere, but humans nearly eradicated the species from the continental United States and Europe by the 1960s (10). Legal protections strengthened during the latter half of the 20th century and wolf populations returned to ten coterminous U.S. states and 28 European countries (11, 12). As of 2019, there are about 5,500 wolves in the United States and 11,000 in Europe. More jurisdictions, such as the state of Colorado, are proposing or already planning reintroductions (13).

Restoring wolves could benefit humans who enjoy seeing wolves in their natural habitat or who value knowing of their existence (14, 15), but in this study we focus on the potential for

¹ Recent research on the cascading benefits generated by recovering sea otters provides a case in point (8, 9). As keystone species, otters provide carbon sequestration benefits via predation on fish that reduce kelp production. Because this benefit is subtle and indirect, it may not resonate with stakeholders who more readily recognize the concrete adverse effects of otters on fish populations that have commercial value.

wolves to generate indirect benefits by controlling overabundant deer populations. Deer populations have surged in the United States, increasing from about 2 to 4 deer per km² in the pre-colonial era (16) to 15 to 50 deer per km² in some areas today (17). Overabundant deer populations affect ecosystems by suppressing forest regeneration, altering the composition of tree and herbaceous plant species, and contributing to the spread of invasive species (18–20). Deer also generate economic costs for humans through deer-vehicle collisions, Lyme disease (which is transmitted through deer ticks), and damage to agriculture, timber products, and landscaping (21).

This study focuses on deer-vehicle collisions (DVCs), which are the largest known economic cost of deer (SI Appendix, Table S1). About one million DVCs occur every year in the United States, causing 29,000 human injuries, 200 human fatalities, and nearly \$10 billion in total economic losses (21, 22).² Europe experiences similar problems, with a lower frequency of collisions with ungulates (such as deer and moose), but a higher rate of fatalities and injuries (23). The problem has worsened over time, with DVCs rapidly increasing since around 1990 (22, 23). If wolves reduce DVCs even modestly, the social and economic benefits could be sizable.

We focus on two channels through which wolves could affect DVCs. The first is changes to deer abundance. We hypothesize that larger wolf populations reduce deer abundance directly through predation (18, 24, 25) which, in turn, reduces DVCs (26–28). The second channel is through changes to deer behavior because wolves create a “landscape of fear” for deer (29–31). Wolves use roads, pipelines, and other linear features as travel corridors, which increases wolves’ travel efficiency and the kill rate of prey near these features (32–36). We hypothesize that wolf presence affects deer movement near these features (34, 37, 38), thereby reducing collision risk for a given number of deer on the landscape. The behavioral effect is important because, through it, wolves can suppress economic damage from deer in ways seasonal human deer hunters cannot.

Our empirical analysis focuses on Wisconsin, where wolves began to recolonize naturally around 1975 (Fig. 1) (39). Wisconsin provides an interesting, if not ideal, case study for several reasons. Foremost, the results from Wisconsin should generalize to other settings where wolves are (or could be) allowed to spread naturally into areas of human communities rather than only into wilderness. Much of the prior research on the effects of wolves in the United States focuses on Isle Royale and Yellowstone National Parks; however, the effect of wolves in National Parks “would have little relevance to most of wolf range because of overriding anthropogenic influences there on wolves, prey, vegetation, and other parts of the food web” (40). By contrast, the spread of wolves across Wisconsin was natural and unimpeded by wildlife managers. As a result, contiguous wolf range in Wisconsin spans a human-dominated landscape more than six times the size of Isle Royale and Yellowstone combined (41). Second, the wolf population in Wisconsin is likely close to ecological carrying capacity (42, 43), which suggests the wolf effects we measure represent a long-term steady state rather than only transitional effects. Finally, if the benefits of a reduction in DVCs outweigh the costs of wolf predation in Wisconsin, then there may be economic justification for allowing wolves to expand not just in the state, but also potentially in other states that have suitable wolf habitat and high DVCs (e.g., the northeastern United States) (44). Although verified predation on livestock and pets in Wisconsin is costly (recently, about \$174,000 per year, on average) (45), DVC losses are orders of magnitude larger (almost \$200 million per year, based on the national average cost per DVC) (22, 46).

² All dollar estimates referenced in this study are converted to 2019 U.S. dollars using the Bureau of Labor Statistics’ Consumer Price Index.

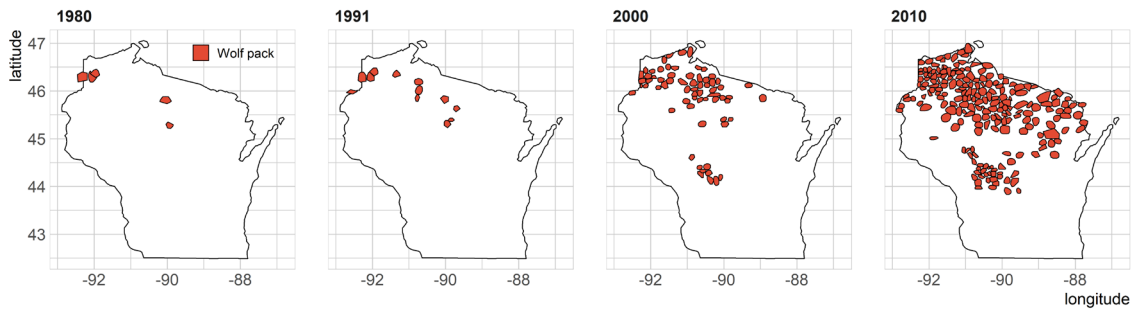


Fig. 1. Wolf packs spread across Wisconsin between 1980 and 2010. Wolf packs are concentrated in the forested areas in the northern and central parts of the state. The rest of the state is predominantly non-forested agricultural or urban areas.

Results

Testing for the effects of wolves requires spatial-temporal data on DVCs, wolf presence, and deer populations. Our preferred study period begins in 1988 (the earliest available DVC data) and ends in 2010 (the last available data on detailed wolf pack locations due to subsequent wolf harvests) (see Methods). As we explain below, the findings are similar when we include years after 2010.

Motivating visualizations

Two differential trends across counties where wolves eventually colonized (wolf counties) and counties where wolves never colonized (non-wolf counties) support our hypotheses regarding wolf effects on DVCs (Fig. 2A-D). First, deer population trends began to diverge in the 1990s, with the populations growing in non-wolf counties and stabilizing in wolf counties. This pattern suggests wolves may be suppressing the trend towards deer overabundance. Second, DVCs were stable in wolf counties but grew in non-wolf counties. By contrast, vehicle collisions not caused by deer (“non-DVCs”) declined at similar rates in areas with or without wolves, suggesting the proportion of vehicle collisions involving deer grew faster in non-wolf counties.

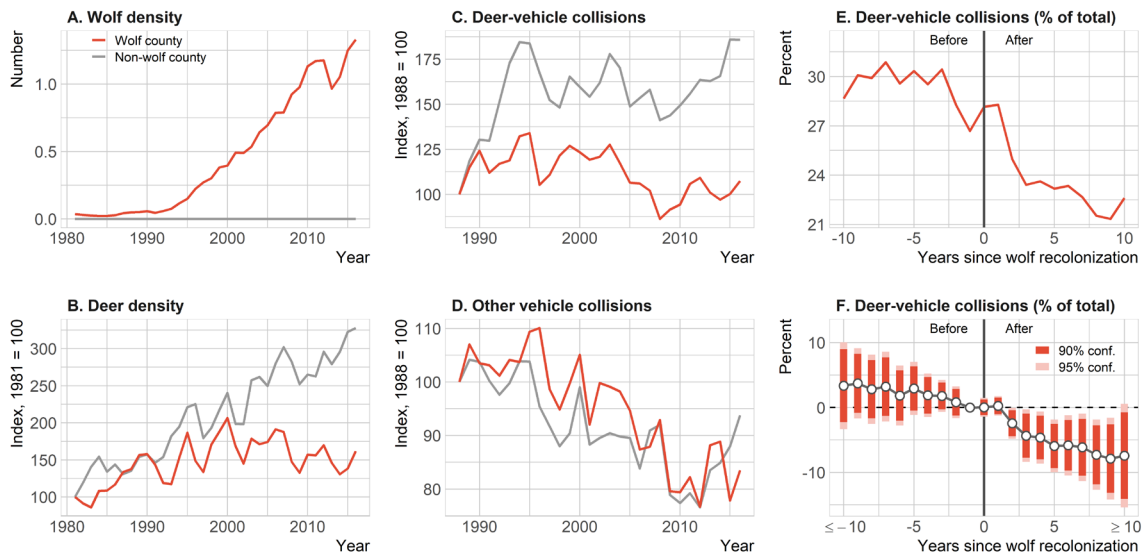


Fig. 2. Trends in wolf abundance, deer abundance, and roadway collisions. (A) Winter wolf population per 100 km² of deer range. Deer range is defined as permanent cover at least 4 ha in

size in 1993, the only year available. (B) Prehunt deer population per km² of deer range. (C) Vehicle collisions caused by deer. (D) Vehicle collisions not caused by deer. (E) Percent of vehicle collisions caused by deer before and after wolves enter a county, 1988 to 2010. (F) Percent of vehicle collisions caused by deer regressed on indicator variables for years since wolf recolonization, county, and year, 1988 to 2010. SEs clustered at the county level. (E and F) Data exclude three counties with wolf exit at some point during the period.

To further motivate the potential effect of wolves, we examine the share of vehicle collisions involving deer before and after wolf entry in each county (Fig. 2E-F). Wolves entered 29 counties by 2010. To visualize the effect of wolf entry, we normalize the year when wolves were first recorded as time $t = 0$ for each county. We exclude three counties for which wolves exit and re-enter at some point during the study period to avoid assumptions about which entry year(s) should be set to $t = 0$. Then we plot the proportion of collisions attributable to deer during each of the 10 years before and after wolf entry (Fig. 2E). The plot reveals a sharp decline in DVC shares that begins in $t = -1$. Although this is one year earlier than expected, the plot is consistent with evidence that dispersing wolves traveling outside heavily forested areas are not immediately detected by the state management agency (43). Analysis of these data demonstrates that, on average, collisions attributable to deer dropped significantly after wolf arrival ($t \geq 0$), relative to years before arrival ($t < 0$) (Fig. 2F). The decline in DVC shares stabilized about 5 years after reintroduction, which is suggestive evidence that deer behavior eventually adapts to wolf presence.

Net effect

To quantify the effects of wolves on DVCs, we apply panel data econometric techniques with 63 counties (29 wolf counties and 34 non-wolf counties) and 22 years (1988-2010). We regress DVCs on annual measures of wolf presence, weather conditions, non-DVCs, and deer populations. All models include fixed effects for each county to control for time invariant characteristics (such as relatively fixed differences in road characteristics across counties) and year-specific effects to control for common, statewide changes over time (such as improvements in vehicle safety or changes in fuel prices).

The preferred model uses a triple-differences approach commonly employed in applied econometrics (47, 48). Applied to our setting, this technique measures whether there is a change in the frequency of DVCs before versus after wolf entry (first difference) in wolf versus non-wolf counties (second difference). Then, it compares whether this “difference-in-differences” is larger for DVCs than for non-DVCs (third difference). For example, if DVCs decline by 30% when wolves enter a county and stay the same in all other counties, then a difference-in-differences estimate would suggest that wolf presence was responsible for a 30% reduction in DVCs. However, if non-DVCs declined by, say, 10% more in non-wolf counties when compared to wolf counties, then the DVC reduction attributed to wolves by the model is 20%. By subtracting the third difference, the model purges the effects of other confounding changes in roadway safety in a specific county that may have occurred at the same time as wolf expansion (e.g. a change in speed limits, road use, driver composition, or road quality). Intuitively, the triple-differences model allows us to control for all time varying factors that affect roadway safety in general but that, unlike wolves, would not specifically affect deer-vehicle collisions.

We first estimate the net effect of wolves on DVCs through changes in both deer population and deer behavior (Fig. 3A). The preferred model shows that wolf presence reduces DVCs by 23.7% for the average county, on net (model 1). The estimated net effect is robust to changes in model specification as indicated by the similar results for Model 2-5. Model 2 excludes the three counties for which wolves exit at some point during the study period to avoid assumptions about the persistence of the effect of past wolf presence. Model 3 excludes the 13 non-wolf counties on the boundary of wolf-counties to rule out possible spatial spillover effects of wolves on neighbor counties. Model 4 flexibly allows wolf counties to have different nonparametric time trends (separate year effects) than non-wolf counties to further allow unobserved annual factors to affect each region differently. Model 5 extends the time series to

2016 rather than concluding in 2010. Collectively, the results suggest that wolves reduced DVCs by 17.9% to 27.3% and that specification choice has little effect on the point estimate. These results are also robust to the inclusion of 8 counties with suspected data quality issues (Materials and Methods and SI Appendix, Fig. S2A).

The net effect of wolf presence on DVCs is economically significant. For the average county with wolves present, a 23.7% reduction in DVCs translates to 38 fewer DVCs per year (Table 1, column 1). Based on the national average loss per DVC (\$9,960) (22), the presence of wolves leads to more than a \$375,000 reduction in DVC losses per county per year; this translates to a savings of about \$11 per person or \$808 per million vehicle miles traveled in 2010. Across the 29 wolf counties, these savings generate a \$10.9 million dollar aggregate reduction in DVC losses each year.

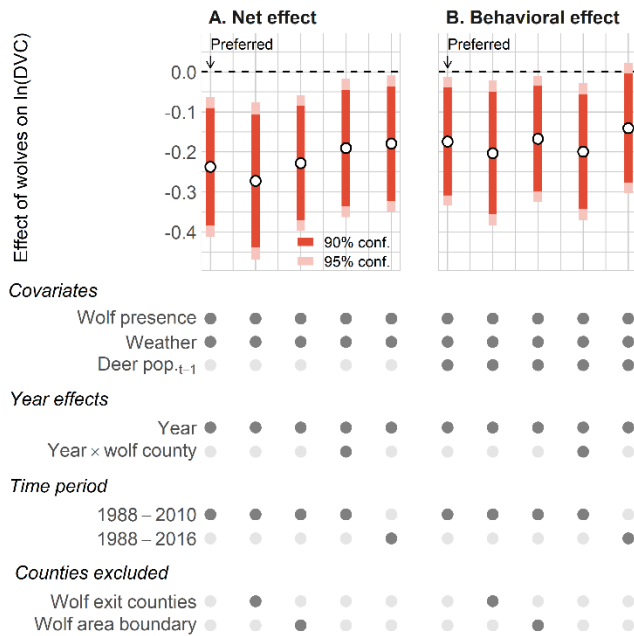


Fig. 3. Wolf presence reduces the frequency of DVCs through changes in deer abundance and behavior. (A) The “net effect” occurs through changes in both mechanisms. (B) The “behavioral effect” occurs through changes in deer behavior only. Subtracting B from A provides the “population effect,” which occurs through changes in deer abundance only. (A and B) Model 2 excludes three counties with wolf exit at some point during the period. Model 3 excludes 13 counties that never have wolf presence (non-wolf counties) on the boundary of counties that have wolf presence for at least 1 year during the study period (wolf counties). SEs clustered at the county level.

Table 1. Wolf presence reduces the economic losses from DVCs

	(1)	(2)	(3)
	Net effect	Behavioral effect	Population effect
Change in DVCs			
Percent	-23.7%	-17.4%	-6.3%
Number	-38	-28	-10
Economic value	\$375,972	\$275,715	\$100,257

Notes: The wolf effect operates through changes in: (1) both deer population and deer behavior, (2) deer behavior only, and (3) deer population only.

Decomposed effect

The net effect of wolves on DVCs embeds a population and behavioral effect. Decomposing the effect is useful for determining whether or not human deer hunting might be a good substitute for wolves in reducing DVCs. If most of the effect on DVCs comes from reductions in deer abundance, then deer hunters could reproduce the benefits of wolves (conditional on sufficient hunting participation). Currently, recreational hunting is the main source of deer population control in the United States (49). If the effect mainly results from changes in deer behavior, then wolves are serving a unique ecological function that seasonal human deer hunters cannot replace.

Controlling for deer abundance in the DVC model separates the behavioral effect of wolves from the population effect (Fig. 3B). This set of models does so by estimating the effect of wolf presence on DVCs conditional on the number of deer on the landscape, thereby purging any effect that wolves have on DVCs through their influence on deer abundance; any remaining effect depends on how wolves change deer behavior. The results suggest that wolf presence reduces DVCs primarily through a behavioral effect; in the preferred model, DVCs fall by 17.4% for the average county (model 1). As with the net effect, the point estimate for the behavioral effect is essentially unchanged across a range of robustness checks (Fig. 3B and SI Appendix, Fig. S2B). For the average county with wolves present, this percentage reduction translates to 28 fewer DVCs per year (Table 1, column 2). The change is valued at about \$276,000 per county per year or \$8 million per year statewide.

The difference between the net effect and the behavioral effect represents the percentage decline in DVCs attributable to the influence of wolf presence on deer populations (or the “population effect”) (Table 1, column 3). Based on this method, wolf presence reduces DVCs by 6.3% through changes in deer abundance. For the average county with wolves present, this translates to 10 fewer DVCs per year, about one-quarter the size of the behavioral effect.

Predator-prey theory

To further evaluate the dominance of the behavioral effect, we estimate the reduction in deer abundance that would be needed to generate the net effect of wolves on DVCs through the population channel only. Then, we compare this benchmark to the estimated effect of wolves on deer abundance from three predator-prey models (Supplementary Information). A 1% decrease in deer abundance leads to a 0.32% reduction in DVCs in the preferred DVC model (SI Appendix, Fig. S1). Therefore, wolf presence would have to reduce deer abundance by 74% to generate the net effect of wolf presence on DVCs if wolf presence had no behavioral effect (i.e. $0.32 \times 74.2 = 23.7$). Equivalently, given the average number of wolves present in a county (14.5), each wolf would need to reduce the deer population by about 5% on average (i.e. $74.2 / 14.5 = 5.1$). However, the predator-prey models show that each wolf reduces the deer population by no more than 1.1% (SI Appendix, Table S2). Overall, the predator-prey models confirm that the effect of wolves on deer abundance is much too small to explain the large effect of wolves on DVCs.

Wolf presence versus wolf abundance

Does the effect of wolves on DVCs depend on the number of wolves in a county? To examine this question, we add to the triple-differences model the annual estimate of wolf populations in addition to the indicator for wolf presence (SI Appendix, Fig. S3A-B). Results indicate that simultaneous inclusion of both measures leaves the effect of wolf presence essentially unchanged (compare model 2, coefficient 1 to the preferred model). The effect of wolf population is statistically insignificant in this model (model 2, coefficient 2), as well as a model that excludes wolf presence (model 3). We interpret these results as suggestive evidence that wolf presence is sufficient to induce behavioral changes in deer that reduce DVCs.

Discussion

Ecologists and economists alike recognize the potential for apex predators to deliver cascading economic benefits to humans, but providing concrete evidence has proven difficult. Unfortunately, the benefits of predators usually become salient only after they are removed from an ecosystem. This was the case in Kern County California, for example, where government programs sought to protect livestock and pet owners by decimating skunks, foxes, badgers, weasels, snakes, owls, ravens, and hawks. Only after the program was it clear that these efforts facilitated the proliferation of mice populations. The mice damaged the local industries that were supposed to benefit from the predator control programs (50).³

Our study suggests that systematic elimination of wolves from North America has also caused unintended damages. Wolves appear to induce an economically significant reduction in the economic losses associated with DVCs. Wolf presence reduces DVC losses by an average of \$375,000 per county per year and by \$10.9 million dollars per year in aggregate across the 29 wolf counties. As a point of comparison, the state paid \$3.1 million in compensation to individuals for verified deaths or injuries caused by wolves of livestock, hunting dogs, and pets between 1985 and 2019, or an average of \$174,000 per year over the last five years (45).⁴ The economic benefit of reduced DVCs exceeds the economic costs of verified wolf predation by a ratio of 63 to 1. This ratio is relevant because economics matters for listing, delisting, and management decisions for endangered species, if only as implicit considerations (56, 57).

The beneficial reduction in DVCs concentrates in rural areas where livestock predation also occurs. This finding may help dampen political polarization around wolf reintroduction that generally pits rural and urban voters against one another, as was the case with the November 2020 vote on wolf reintroduction in Colorado (58). In contrast to studies that focus on the economic benefits to humans in urban areas who value knowing of wolf existence (14), our study's focus may be of more practical use to policy makers trying to balance competing constituent interests when setting policy.

Furthermore, the findings imply that wolves may provide a solution to an intractable traffic safety engineering problem, namely that traditional DVC mitigation measures often are not scalable over large regions. Out of more than 40 types of traditional mitigation measures, only 13 are effective at reducing collisions with large ungulates (22).⁵ Of these strategies, most are cost-effective (meaning, the present value of benefits exceeds the present value of costs), only if deployed in DVC hotspots (59); hence, information and funding constraints limit the widespread use of mitigation strategies not involving wolves. Newer technologies, such as machine learning models that identify the presence of deer on roads (60, 61) and rear-facing illumination systems that make vehicles more visible to deer (62), have shown promise in reducing DVCs. It remains unclear how quickly and widely these technologies can be deployed.

³ In the jargon of economics, the excessive damage resulting from predator removal is an "ecosystem externality," given the effects of this activity on third parties (50).

⁴ Compensation programs are the most common metric used to estimate the economic costs of wolf predation on livestock, hunting dogs, and pets (51); however, this measure likely under-estimates the true effect. For example, it does not account for some direct but unverifiable losses, such as missing livestock (rules about compensation for missing calves vary over time in Wisconsin). It also does not account for indirect effects of stress on the growth, reproduction, and health of livestock (52). These indirect effects are difficult to estimate but may be as large or larger than direct costs (53). Compensation programs also do not account for the time and money spent by producers to defend their herds and reduce the consequences of attacks (54). Nor does it account for the distress of livestock and pet owners caused by lost animals and their willingness-to-pay to avoid wolf predation (55). Although these and other indirect costs of wolves are difficult to measure, they should be considered in management decision-making.

⁵ Effective measures include seasonal wildlife warning signs; vegetation removal; fence, gap, and crosswalk; population culling; relocation; anti-fertility treatment; fence; fence, underpass, and jump-out; fence, under- and over-pass, and jump-out; animal detection system (ADS); fence, gap, and ADS; elevated roadway; and road tunnel. Effectiveness ranged from 26% to 100% (59).

In any case, the changes in deer behavior induced by wolf presence could have additional positive effects beyond reducing DVCs. Deer regularly forage on the edges of agricultural fields (16, 63), which tend to be near roads. When wolves are present, deer may retreat from the fields to wooded areas with better cover (64, 65). This behavioral change could be a substantial benefit to farmers, as deer cause 90% of all wildlife damage to agriculture in Wisconsin (66).⁶ Wolves may also reduce the incidence of Lyme disease by changing the ecology of small mammal hosts (68).⁷ Lyme disease has annual economic costs of \$4.4 billion to \$7.9 billion (69–71). Wolves also can mitigate the harmful effects of overabundant deer on forest ecosystems (3–5).

Importantly, our study implies that wolves regulate deer populations and behavior in ways that a waning number of human deer hunters cannot. In many areas, deer are considered overabundant (18, 19), and managers have struggled to incentivize recreational hunters to reduce deer numbers to a level that would reduce the harmful ecological effects of deer (72–74). A sharp decline in deer hunting participation in recent years will make it more challenging to rely on recreational hunters (75, 76). More fundamentally, deer hunters have limited ability to change deer behavior. Although deer avoid roads when deer hunters are present (77, 78), most hunting activity (and, therefore, any potential effect on DVCs) is compressed into a few weeks each year. In contrast, wolves alter deer behavior year-round. This explains why most (73%) of the estimated effect of wolves on DVC reductions is attributable to their effects on deer behavior.

The finding that wolves reduce DVCs primarily by changing deer behavior rather than by reducing deer abundance is likely good news for policy makers. It implies they do not need to choose between a \$20.6 billion nationwide recreational deer hunting industry (79) and DVC benefits from wolves. At least in Wisconsin, it seems that wolves and deer hunters can co-exist, with safer roadways. This is consistent with research from Europe, suggesting the potential for successful coexistence with wolves on that continent as well (11).

Materials and Methods

Literature

Wolf-deer interactions

A growing body of literature finds that the recolonization of wolves precedes dramatic ecological changes in forested areas, such as increases in aspen, willow, and cottonwood tree recruitment (4, 80–82), forb and shrub species richness (5), and balsam fir growth (3). The theory is that prey are more cautious in ecosystems with wolves, which leads to less time spent in any one place and in areas with high predation risk. This avoidance behavior reduces herbivory pressure on preferred food sources, which in turn affects other aspects of the ecosystem, a process known as a trophic cascade. Although debate remains about whether wolves caused these observed ecosystem changes or were merely correlated with them (40, 83), there is strong circumstantial evidence that wolves have important and quantifiable impacts on a variety of ecological outcomes.

There are direct and indirect channels through which wolves could affect deer-vehicle collisions (DVCs) through a trophic cascade. The direct mechanism is changes to deer abundance. It is intuitive that wolf predation would decrease deer abundance, but in reality this relationship is complicated and difficult to identify empirically. There is some evidence that wolves

⁶ Deer-inflicted crop damage was last valued at \$74 million to \$92 million per year in 1996, when deer populations were substantially lower than they are today (67).

⁷ The theorized impact of wolves on Lyme disease operates through changes in predator-prey interactions (68). In short, the absence of wolves has led to rapid growth in coyote (*Canis latrans*) populations. Coyotes suppress the abundance of several small mammal predators, such as red fox (*Vulpes vulpes*), which leads to increases in small mammal populations. These small mammals are thought to infect the majority of ticks that spread Lyme disease.

can suppress elk (*Cervus canadensis*), deer (*Odocoileus* spp.), and moose (*Alces alces*) populations (18, 24, 25, 84), but this effect may be mediated by many factors, such as the relative levels of predator and prey, forage availability, and weather (85–87). The effect of a reduction in deer abundance on DVCs could be economically significant. The only prior study of which we are aware that examines the impact of carnivores on DVCs predicts, using theoretical simulations, that the reintroduction of cougars (*Puma concolor*) to the eastern United States could reduce DVC losses by \$2.3 billion over 30 years through reductions in deer density (28).

The indirect mechanism is changes to deer behavior. Areas with high road density are generally poor habitats for wolves (88–92); however, wolves also use roads, pipelines, and other linear features as travel corridors, which increases wolves' travel efficiency and the kill rate of prey near these features (32–36). As such, prey may avoid these features when wolves are present (34, 37, 38). It makes sense that these behavioral changes also would cause deer to avoid roads, thereby reducing DVCs; however, to our knowledge this effect has yet to be measured. The behavioral effect is particularly important because it suggests that wolves suppress economic damage from deer in a way that hunters cannot.

Deer-vehicle collisions

Economists have investigated many factors that influence traffic safety, such as mandatory seat-belt laws and speed limits, drunk driving, fuel economy standards, vehicle size, and Daylight Saving Time (93–97). Most prior economic studies focus on fatal collisions, which are critical to understand because mitigating such collisions prevents the loss of human life; but, fatal collisions only account for 0.2% of all collisions (98). Research on the remaining 99+% of collisions could illuminate strategies to minimize the impacts of crashes on a large segment of the population. Deer-vehicle collisions (DVCs) constitute 7% of all vehicle collisions and 24% of single-vehicle collisions (22, 99).

Three categories of variables describe the main determinants of DVCs in the literature: exposure, road characteristics, and land cover (26, 100–116). Intuitively, both larger deer populations and higher traffic volume tend to increase DVCs by increasing the probability a driver will encounter (or be exposed to) a deer. Road characteristics, such as curvature, steepness, and number of lanes, affect DVCs by influencing driver reaction times or changing deer movement patterns.⁸ The effects of land cover on DVCs is mixed, with some studies finding positive impacts and some finding negative impacts of agricultural, forested, and urban areas. Landscape diversity and fragmentation appear to be more important and tend to increase DVCs; however, these variables may be proxies for deer abundance because deer tend to prefer landscapes that are more heterogeneous.

The effects of weather on non-deer vehicle collisions is well-established (117–127), but only three animal-vehicle collisions (AVC) studies of which we are aware examine weather (128–130). Collectively, these studies demonstrate that weather affects both deer and driver behavior, sometimes in opposing ways, and the impacts of weather on DVCs is context specific and poorly understood.⁹

⁸ Bridges, drainage ditches, jersey barriers, gullies, and right-of-way fencing (less than 1.5 m or 5 ft tall) are associated with higher DVCs, as are roads with higher curvature, more than two lanes, or short or obstructed sight lines. Roads with steeper or groomed rights-of-way tend to decrease DVCs.

⁹ There is a consensus that rain increases the frequency of vehicle collisions by causing slippery conditions and reduced visibility, but no AVC study includes rain. Snow can either increase or decrease vehicle collisions, depending on driver defensive behaviors. Snow also increases both moose-vehicle collisions (MVC) and AVCs (128, 129), but the effect on DVCs is unknown. Cold temperatures increase vehicle collisions, but may decrease fatal collisions (118, 123, 125). Cold temperatures increase MVCs because moose increase movement at low temperatures (129); the effect on DVCs is unknown. High temperatures (greater than 26.7 °C or 80 °F) increase vehicle collisions, likely by increasing driver irritability and reaction times. Summer temperatures did not affect MVCs in Norway (129); however, deer decrease activity or seek thermal cover at these temperatures (131), which would tend to reduce DVCs. Positive phases of the North Atlantic Oscillation increase DVCs in spring in Norway, but the relative contributions of temperature, precipitation, and snow depth in this relationship is unknown (130).

Most previous research on DVCs focuses on identifying the characteristics of “hotspots” or short stretches of roads with especially frequent crashes (132). Understanding these characteristics helps traffic safety engineers target mitigation strategies (e.g. fences and deer warning signs) and design safer roads. Risk maps also shed light on the biological mechanisms causing DVCs, and they may have long-lasting predictive power (133, 134). However, it is difficult to extrapolate hotspot risk maps regionally. As a result, there is a gap in understanding how to prevent DVCs cost-effectively over large regions, especially in the presence of an apex predator.

As described in more detail below, we estimate the relationship between wolves, deer abundance, and DVCs at a broader spatial scale – counties in Wisconsin. This is important because location-specific interventions could displace DVCs to different areas, and research focused on specific locations will fail to encompass these displacement effects. In addition, it is difficult to estimate deer population in a small roadside buffer; using a larger unit of analysis allows us to better measure deer population. A handful of previous studies measure how county-level characteristics affect the frequency of ungulate-vehicle collisions (27, 129, 130, 135–138). The findings from these studies are generally in line with the hotspot literature.

The existing county-level studies provide a foundation for identifying determinants of DVCs at a broad geographic scale; however, several data and methodological limitations may introduce undesirable properties of coefficient estimates. Small sample sizes (N=9 (135) and N=5 (130)), imprecisely measured covariates (27, 129, 130, 137, 138), replication of data values across years as proxies for missing values (27), and inclusion of few covariates (130, 137) could all contribute to biased or inconsistent coefficient estimates. Lastly, most of these studies pooled data across time for spatial analysis and across counties for time series analysis. In contrast, we develop a unique panel dataset covering 72 counties over 28 years, which allows us to control for a bundle of county-level characteristics that might affect DVCs using modern econometric techniques; none of the existing studies use the panel data methods employed here.

Data

We assembled an annual, county-level data set for 1979 through 2016 that includes DVCs, deer and wolf populations, and weather. The study period is limited by data availability before 1988 for DVC’s and after 2010 for detailed wolf pack locations. Although we can estimate county-level wolf presence data for 2010 to 2016, the source data are collected on a different basis during that period; this change increases measurement error. Below we describe each data set.

Deer-vehicle collisions

The Wisconsin Department of Transportation provided county-level, police-reported DVCs for 1988 through 2013. The Wisconsin Traffic Operations and Safety Laboratory at the University of Wisconsin-Madison provided crash-level data for 1994 through 2016. All overlapping years are identical.

We exclude 8 of 72 counties because police in these counties apparently stopped attending DVC crash scenes during the study period, except when the crash causes a road hazard, human injury, or fatality, based on discrete structural breaks in the data. The main results are robust, however, to inclusion of these counties, as a comparison of Fig. 3 with SI Appendix Fig. S2 makes clear.

We also exclude Menominee County, which is coterminous with the boundaries of the Menominee Indian Reservation. The state of Wisconsin does not have reliable data on the variables of interest for this region.

Deer population

The Wisconsin Department of Natural Resources (WDNR) provided posthunt white-tailed deer population by Deer Management Unit (DMU) for 1981 through 2016. Historically, Wisconsin’s DMU boundaries followed roads or other natural features that would be easy for

hunters to identify in the field. In 2013, the WDNR revised DMUs to follow primarily county boundaries. At that time, they estimated deer population for the new DMUs for 2002 onward by distributing each old DMU's deer population to the new DMUs it overlaps based on an area-weighted measure. We replicated this procedure for 1981 onward. The procedure assumes that deer are distributed uniformly across the deer range in each DMU. Deer range includes all permanent cover—forest, woodlot, brush-covered land, or marsh—at least four hectares in size in 1993, the only year available. We also converted population estimates from posthunt to prehunt using the WDNR's formula for this conversion (139), namely $\text{prehunt population} = \text{posthunt population} + \text{harvest} \times 1.15$.

Wolf population

WDNR provided mid-winter wolf pack locations and populations for the winter spanning 1979 and 1980 through the winter spanning 2010 and 2011. This is the last year for which WDNR publishes individual wolf pack locations, likely due to legislative changes. The federal government temporarily delisted wolves from the Endangered Species Act (ESA) on January 27, 2012 (43). The state legislature designated the gray wolf as a game species and authorized wolf harvest on April 2, 2012 (140). Then, WDNR stopped publishing wolf pack locations in the September, 2012 data release, which covers wolf status during the winter spanning 2011 and 2012 (141). Beginning with that release, WDNR provides only the number of wolves in survey blocks. These data sets include some lone wolves; however, WDNR's monitoring methods focus on wolf packs, and lone wolves likely remain undercounted (43).

In most models, the main variable of interest is a dummy variable for wolf presence. The preferred models define county level wolf presence based on wolf pack locations. If at least one wolf pack's territory crosses a county boundary, we define that county as having wolf presence. We also extend the times series using the wolf survey block data as a robustness check; namely, for 2011 to 2016, if any populated wolf survey block overlaps a county boundary, we define that county as having wolf presence. This procedure implicitly assumes that wolves range across the entire survey block.

Finally, we calculate wolf population using the same procedures as wolf presence, except we assign wolf pack or survey block population to counties using a spatially-weighted measure. The population measure allows us to test the hypothesis that wolf presence alone is sufficient to alter deer behavior, even though higher wolf abundance may be required to reduce deer abundance.

Weather

Daily minimum temperature, maximum temperature, and precipitation (rain plus melted snow) are available online as 4-kilometer grid raster files from the Parameter-elevation Regressions on Independent Slopes Model (PRISM) Climate Group at Oregon State University beginning in January 1, 1981. We calculate spatially-weighted averages for each variable, county, and day. Then, we calculate total annual precipitation (cm), the number of days with a minimum temperature below 0 °C (32 °F), and the number of days with a maximum temperature above 26.7 °C (80 °F), for each county and year. The rationale for using these weather covariates is described in the literature review above.

Conceptual Models and Estimation Methods

Net effect

As a starting point for identifying the net impact of wolves on DVCs through changes in deer abundance and deer behavior, we begin with a difference-in-differences (fixed-effects) model that is commonly used in modern econometrics (47, 48). Applied to our setting, this technique measures whether there is a change in the frequency of DVCs before versus after wolf entry (first difference) in wolf versus non-wolf counties (second difference). Rather than trying to include measures for every factor that could affect DVCs within a particular county, this model uses non-parametric methods – indicator variables for each county – to control for all time-

invariant factors that could affect average differences in DVC rates across counties (e.g., differences in average roadway curvature across counties). The standard model also uses non-parametric methods – indicator variables for each year – to control for time-variant factors that could affect DVCs across the entire state of Wisconsin (e.g., gasoline prices, improvements in vehicle safety, or annual weather fluctuations). The standard model is as follows:

$$\ln(DVC_{it}) = \alpha WolfPresence_{it} + \beta'W_{it} + \theta_i + \delta_t + u_{it} \quad (1)$$

where DVC_{it} is the number of vehicle collisions in county i at time t , $WolfPresence_{it}$ is a dummy variable indicating whether a wolf pack is present, θ_i is a vector of county effects, δ_t is a vector of year effects, and u_{it} is the idiosyncratic error. Although δ_t accounts for annual differences in weather severity across Wisconsin, we also include the matrix W_{it} to account for annual differences in severity within each county. W_{it} includes the natural log transformation of the following covariates: total annual precipitation (cm), the number of days with a minimum temperature below 0 °C (32 °F), and the number of days with a maximum temperature above 26.7 °C (80 °F). All weather variables are always greater than zero. The rationale for using these weather covariates is described in the literature review above. We hypothesize that α , the net effect of wolf pack presence on the frequency of DVCs, is negative. Note that the model estimates the effects of deer and wolves on DVCs net of drivers adapting to changing risks of collision. As such, the estimates will not reflect the economic costs of avoidance behaviors (e.g., driving at different times or taking different routes to avoid deer) and will therefore understate the full economic benefits of any wolf-induced reduction in DVCs.

To summarize, the empirical strategy in Eq. (1) is based on within-county variation in DVCs over time based on wolf entry and exit over time within a given county. For the most part, this variation is based on initial entry to the county because wolves tend to stay in a county permanently. However, there are three counties for which wolf packs entered and exited multiple times. These are likely cases where wolf pack territories are straddling the county boundary. The model captures these dynamics. If wolves exit a county, the wolf presence dummy variable switches to zero; it switches back to one if wolves re-enter.

The main limitation is that the estimation strategy in Eq. (1) does not control for time varying factors that might have differentially affected roadway safety across counties within Wisconsin for reasons unrelated to deer and wolves. Such factors could include changes in the relative amount of roadway traffic, changes in the relative maintenance and quality of roads, or changes in the relative demographic composition of drivers. We address this limitation by utilizing the annual county-level data on non-deer vehicle collisions (non-DVCs). Our assumption is that changes in non-DVCs will reflect time-varying changes in road safety conditions that are unrelated to deer abundance and wolf presence.

Controlling for changes in non-DVCs builds our estimated model into the triple-differences model given by Eq. (2). Triple-differences estimators can “generate a more convincing set of results than a traditional [difference-in-differences] analysis” (47). They are also widely used in the econometrics literature, appearing in over 900 peer-reviewed papers since 2000 (48). The model is as follows

$$\ln(VC_{it}) = \alpha_{11}WolfPresence_{it} + \beta_{11}'W_{it} + \theta_{11i} + \delta_{11t} + (\alpha_{12}WolfPresence_{it} + \beta_{12}'W_{it} + \theta_{12i} + \delta_{12t}) \times DVCDummy_{it} + u_{it} \quad (2)$$

where VC_{it} is the number of vehicle collisions (of any type) in county i during year t and $DVCDummy_{it}$ is an indicator variable indicating whether a collision was caused by deer. By interacting the weather matrix W_{it} with $DVCDummy_{it}$, we allow weather to have a different effect on deer versus non-deer vehicle collisions. This flexible approach allows weather to affect accidents not only through its effects on general roadway safety conditions, but also through its effects on deer behavior and abundance. The difference-in-differences estimate of α_{11} measures changes in vehicle collisions of any type, DVC or non-DVC, before versus after wolf arrival (first difference) in wolf versus non-wolf counties (second difference). We have no hypothesis about the sign of α_{11} because this coefficient estimate will capture any generic changes in roadway safety that correlate

with the timing of wolf arrival and presence in a county. We hypothesize that α_{12} is negative. It is the triple-differences coefficient that assesses how the number of DVCs changes relative to non-DVCs before versus after wolf arrival in wolf versus non-wolf counties (the third difference). Intuitively, this model allows us to control for all time varying factors that affect roadway safety in general but that, unlike wolves, would not specifically affect deer-vehicle collisions. For comparison with other models, discussed below, we convert the proportional effect of wolf presence to a level effect by multiplying the point estimate for α_{12} by the annual number of DVCs in the average county in which wolves are present.

We estimate a number of robustness checks to account for other factors that could possibly confound the estimates of how wolf presence affected DVCs. First, we exclude three counties for which wolves exit at some point during the study period to avoid assumptions about the persistence of the effect of past wolf presence. Second, we exclude 13 non-wolf counties on the boundary of wolf-counties to rule out possible spatial spillover effects of wolves on neighbor counties. Third, we use separate year effects for counties that ever get wolves versus those that do not to allow for the possibility that DVCs were trending differently in wolf vs. non-wolf counties before the wolves arrived, even relative to non-DVCs. Finally, we extend the study period through 2016 to examine the impact of timing on the model estimate. For all models, we cluster standard errors at the county level.

One concern could be that wolves are spreading spatially in response to changes in deer populations. However, two studies in Wisconsin fail to find a significant difference in deer density between wolf pack areas and non wolf pack areas. The studies also find that deer density is not a significant predictor of wolf habitat suitability (89, 90). This finding is helpful for causal identification because it suggests there is not reverse causation with wolves responding to deer abundance.¹⁰

Behavioral effect

To measure the effect of wolves on DVCs through changes in deer behavior, we control for deer abundance in Eq. (2). In other words, what is the effect of wolves on the frequency of DVCs, for a given number of deer on the landscape? The model is as follows:

$$\ln(VC_{it}) = \alpha_{21}WolfPresence_{it} + \gamma_{21}\ln(DeerPop_{it-1}) + \beta_{21}'W_{it} + \theta_{21i} + \delta_{21t} + (\alpha_{22}WolfPresence_{it} + \gamma_{22}\ln(DeerPop_{it-1}) + \beta_{22}'W_{it} + \theta_{22i} + \delta_{22t}) \times DVCDummy_{it} + u_{it} \quad (3)$$

where $DeerPop_{it-1}$ is the prehunt deer population for county i during year $t-1$, and the other variables are defined as previously. Deer population is lagged for two reasons: 1) to reduce potential endogeneity between vehicle collisions, which typically result in the death of the deer, and deer populations and 2) because contemporaneous wolf presence cannot affect lagged deer populations. We hypothesize that α_{22} , the effect of wolf presence on the frequency of DVCs through changes in deer behavior, is negative. For comparison with other models, we convert the proportional effect of wolf presence to a level effect by multiplying the point estimate for α_{22} by the annual number of DVCs in the average county with wolves present over the study period. We also estimate the same robustness checks for this model as we did for the net effect model, described above. For all models, we cluster standard errors at the county level.

Population effect

The difference between the net effect and the behavioral effect represents the preferred estimate of wolf presence on DVCs through changes in deer population (or the “population effect”). We calculate this effect by subtracting the point estimate for α_{22} (Eq. (3)) from α_{12} (Eq.

¹⁰ The literature's finding that temporal colonization of wolves across Wisconsin appears to be exogenous to changes in deer abundance and DVCs adds to our confidence that the effects we estimate are causal. We emphasize, however, that our results should be interpreted as an average effect for counties with suitable wolf habitat, rather than an effect that could be projected to counties without suitable wolf habitat. In econometric jargon, we are most credibly estimating an ATT (average treatment effect on the treated counties) rather than an ATE (average treatment effect applicable to all counties). The difference between the two estimates is described in modern econometrics textbooks (notably, 47).

(2)). Note that this effect represents how wolf presence affects DVCs through changes in deer abundance, assuming that deer behavior has simultaneously changed in response to wolves. As before, for comparison with other models, we convert the proportional effect of wolf presence to a level effect by multiplying the point estimate for the population effect by the annual number of DVCs in the average county with wolves present over the study period.

Predator-prey theory

To corroborate the relative magnitudes of the population effect versus the behavioral effect, we estimate the effect of wolves on county-level deer populations using three predator-prey models (Supplementary Information). The first estimate is the simplest (and most naïve). It assumes that one additional wolf reduces the deer population by the average number that it eats, which is 20 deer per year in Wisconsin (142). The naïve model ignores compensatory mortality (i.e., wolves kill some deer that would have died of other causes) and non-consumptive effects of wolves on deer (e.g., effects of energetically costly avoidance behaviors and stress on reproductive success) (52). The second estimate applies a dynamic predator-prey model to predict deer abundance at the county level based on changes in lagged deer abundance, wolf abundance, and winter severity (139, 143, 144). In contrast to the simple estimate, the dynamic model also accounts for indirect effects that wolves have on deer populations, for example, by reducing the reproductive success of deer. Although this type of predator-prey model is standard in the ecological literature, including both lagged deer and lagged wolf abundance in one specification may bias the coefficient estimates, as these two variables are simultaneously determined. As such, the third estimate excludes lagged deer abundance from the predator-prey model. Finally, we compare the ecological models to the implied reduction in deer population needed to reproduce the net effect of wolves from the DVC model, converted to a per wolf basis by dividing the effect of wolf presence by the average number of wolves present.

Wolf presence versus wolf abundance

To examine whether the effect of wolves on DVCs increases with wolf abundance, we test the effects of two wolf data transformations on the net effect model in Eq. (2) and the behavioral effect model in Eq. (3). The first specification includes wolf presence; this is the preferred model, as defined above. The second specification includes both wolf presence and wolf population; this model provides insight on the relative importance of wolf presence versus wolf abundance. The third specification includes wolf population. Comparing model results helps us determine if wolf presence is the main driver of DVC changes.

Economic effects

For all models, we estimate the economic effects of changes in DVCs by multiplying the level change in DVCs by the national average of economic losses per DVC, \$9,960 (22). Economic losses includes vehicle repair costs, towing and law enforcement services, monetary value of the animal, and carcass removal and disposal. They also include “lost earnings, lost household production, medical costs, emergency services, travel delay, vocational rehabilitation, workplace costs, administrative and legal costs, and pain and lost quality of life.”¹¹

Data and code availability

The datasets generated during and/or analyzed during the current study are available in Dryad at <https://doi.org/10.5061/dryad.g4f4qrfp8>; code is available at https://github.com/jennifer-raynor/wolvs_and_DVCs. This paper does not use any custom algorithms.

¹¹ The value of these costs is estimated using “willingness-to-pay cost,” which represents “the costs people are willing to pay for safety improvements to avert a fatality or injury” (145). This measure does not quantify the costs associated with the emotional distress of people and animals involved in DVCs, nor the costs of cultural values impacted by wounded animals (22).

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